

Power-law and exponential tails in a stochastic priority-based model queue

G. Grinstein and R. Linsker

IBM T. J. Watson Research Center, P. O. Box 218, Yorktown Heights, New York 10598, USA

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We derive exact asymptotic results for a stochastic queueing model in which tasks are executed according to a continuous-valued priority. The distribution $P(\tau)$ of the waiting times τ of executed tasks for this model is shown to behave asymptotically as a power law, $P(\tau) \sim \tau^{-3/2}$, when the average rates of task arrival λ and execution μ satisfy $\mu \leq \lambda$ (as was earlier noted empirically). For $\mu > \lambda$, $P(\tau) \sim \tau^{-5/2} \exp[-(\sqrt{\mu} - \sqrt{\lambda})^2 \tau]$.

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I. INTRODUCTION

In an interesting recent series of papers [1–3], Barabási and co-workers analyzed data on activities such as web browsing, library use, and the exchange of letters and e-mail messages. Studying the distributions $P(\tau)$ of waiting times τ between successive events—for example, the time it takes for the recipient of a letter or e-mail message to reply to it—they reported that the distributions exhibited heavy tails consistent with power laws $P(\tau) \sim \tau^{-\alpha}$ over some range of τ . The reported values of α were close to 3/2 for written correspondence and close to 1 for the other activities. These results stand in contrast to much of the large literature on waiting-time distributions for real and model queues, where exponential decays of $P(\tau)$ are typical [4].

Several elementary queueing models were considered in Refs. [1–3] to try to explain the observed behavior. One of these models was the continuous-priority version of a classic model by Cobham [5]—a stochastic queue wherein tasks are selected for execution on the basis of a discrete task priority value. The continuous-priority version of this model was indeed found numerically to exhibit asymptotic power-law behavior of $P(\tau)$, with exponent 3/2, provided the average rate λ of task arrival equals [1–3] or exceeds [2,3] the average rate μ of task execution. We subsequently derived this result analytically [6], as well as the new result [6] $P(\tau) \sim \tau^{-5/2} \exp(-\tau/\tau_0)$ for $\mu > \lambda$, by approximating the number of tasks in the queue by a continuous variable, and thereby mapping the model onto the familiar problem of biased diffusion; here τ_0 is a characteristic time that diverges as $(\mu - \lambda)^{-2}$ as λ approaches μ from below.

In the present paper, we solve the model without approximation, deriving exact asymptotic results for the realistic case in which the number of tasks is an integer [7]. We show that indeed $P(\tau) \sim \tau^{-3/2}$ for $\mu \leq \lambda$, and that $P(\tau) \sim \tau^{-5/2} \exp[-(\sqrt{\mu} - \sqrt{\lambda})^2 \tau]$ for $\mu > \lambda$.

II. THE MODEL: STOCHASTIC, CONTINUOUS-PRIORITY QUEUE

In the continuous-priority version [1] of the Cobham model [5], new tasks, having priority x (with $0 \leq x \leq 1$) chosen randomly from an arbitrary probability distribution $\rho(x)$ (which we make uniform by a simple transformation, without any loss of generality [6]), arrive in the queue with an average rate λ . The highest-priority task in the queue is ex-

ecuted with an average rate μ . As in Ref. [1], we consider the case in which the removal from the queue of a task that has been selected for execution occurs instantaneously (i.e., with zero service time).

III. ANALYSIS

In a manner similar to Ref. [6], the overall probability per unit time, $P(\tau)$, that a given task sits in the queue for a time τ before being executed is conveniently expressed in terms of two quantities: (a) the probability per unit time, $G(n, x, \tau)$, that a given task of priority x , which arrives in the queue at time $t=0$ with exactly n items of higher priority (i.e., larger x) already in the queue, gets executed at time τ , and (b) the probability $\tilde{Q}(n, x)$ of there being exactly n items in the queue with priority greater than x , once a steady state has been achieved [cf. Ref. [8], Eq. (2.10)]:

$$P(\tau) = \sum_{n=0}^{\infty} \int_0^1 dx \tilde{Q}(n, x) G(n, x, \tau). \quad (1)$$

As in [6], let $Q(m, x, t)$ be the probability that at time t there are exactly m tasks with priority greater than x in the queue, for $m=0, 1, 2, \dots$. Then $Q(m, x, t)$ satisfies the following master equations, for $m > 0$ and $m=0$, respectively:

$$\begin{aligned} \partial Q(m, x, t) / \partial t = & aQ(m+1, x, t) + bQ(m-1, x, t) \\ & - (a+b)Q(m, x, t), \end{aligned}$$

$$\partial Q(0, x, t) / \partial t = aQ(1, x, t) - bQ(0, x, t); \quad (2)$$

here $a = \mu$ and $b(x) = \lambda \int_x^1 \rho(z) dz = \lambda(1-x)$ are the respective probabilities per unit time of the number of tasks with priorities greater than x in the queue decreasing by 1 (provided $m \geq 1$) and increasing by 1. (We will generally suppress the explicit x dependence of b in what follows.)

As in Ref. [6], it is simple to show that for $\mu > \lambda$ the steady-state solution of Eqs. (2) achieved as $t \rightarrow \infty$, where $\partial Q(m, x, t) / \partial t$ vanishes for all m , is

$$\tilde{Q}(m, x) = (1 - b/a)(b/a)^m. \quad (3)$$

Since (again, as in [6]), $b(0)$ approaches a as λ approaches μ from below, the distribution $\tilde{Q}(m, 0)$ becomes uniform in m , and the mean number of tasks in the queue in steady state, $\langle m(x=0) \rangle$, diverges as $1/(\mu - \lambda)$. Thus, though in this strict

sense the steady-state distribution is ill defined for $\lambda = \mu$ [4,8], in fact the queue does have well-defined steady-state properties since, for $\lambda = \mu$, the mean number of tasks with priorities greater than x is $\langle m(x) \rangle = (1/x - 1)$ and thus remains finite for any $x > 0$. (The case $\mu < \lambda$, shown in [6] to be closely related to the case $\mu = \lambda$, will be summarized at the end of the paper.)

We continue with the analysis, restricting ourselves to $\mu \geq \lambda$ for now. To compute $G(n, x, \tau)$, consider $p(m, x, t)$, the probability that at time t there are exactly m tasks with priority greater than or equal to x in the queue, and that these tasks include the task with priority x that arrived at time $t = 0$; here m is a positive integer. For $m > 1$ and $m = 1$, respectively, $p(m, x, t)$ satisfies master equations similar to those for $Q(m, x, t)$:

$$\begin{aligned} \partial p(m, x, t) / \partial t &= ap(m+1, x, t) + bp(m-1, x, t) \\ &\quad - (a+b)p(m, x, t), \\ \partial p(1, x, t) / \partial t &= ap(2, x, t) - (a+b)p(1, x, t). \end{aligned} \quad (4)$$

Since there are exactly n tasks with priority greater than x at $t=0$, the initial condition on these equations is $p(m, x, t=0) = \delta_{m, n+1}$.

The only difference between the equations for Q and for p is at the boundary: It is easy to see that the total probability $Q_{\text{tot}}(x, t) \equiv \sum_{m=0}^{\infty} Q(m, x, t)$ has a vanishing time derivative $\dot{Q}_{\text{tot}}(x, t) = 0$, whereas the derivative $\dot{p}_{\text{tot}}(x, t)$ of the probability $p_{\text{tot}}(x, t) \equiv \sum_{m=1}^{\infty} p(m, x, t)$ satisfies $\dot{p}_{\text{tot}}(x, t) = -ap(1, x, t)$. This reflects the fact that $m=0$ is an absorbing boundary for $p(m, x, t)$: When $m=1$, only the task with priority x that arrived in the queue at $t=0$ is present. If that task is executed, then the process of emptying the queue of tasks whose priorities equal or exceed x is complete, which implies that $G(n, x, \tau) = -\dot{p}_{\text{tot}}(x, \tau) = ap(1, x, \tau)$.

Equations (4), which represent a random walk with a drift velocity (see [9] and [6]), are solved by standard Laplace transform methods (e.g., [9]), the chain of equations for the Laplace transform $\tilde{p}(m, x, s)$ of $p(m, x, t)$ taking the forms, for $m > 1$ and $m = 1$, respectively,

$$\begin{aligned} (s+a+b)\tilde{p}(m, x, s) &= a\tilde{p}(m+1, x, s) + b\tilde{p}(m-1, x, s) + \delta_{m, n+1}, \\ (s+a+b)\tilde{p}(1, x, s) &= a\tilde{p}(2, x, s) + \delta_{1, n+1}. \end{aligned} \quad (5)$$

It is easy to verify that the solution of these equations is given by $\tilde{p}(m, x, s) = c_m \tilde{p}(1, x, s) + d_m$, where $c_m \equiv (\beta_+^m - \beta_-^m) / (\beta_+ - \beta_-)$ and $d_m \equiv \theta(m-n-1)(\beta_-^{m-n-1} - \beta_+^{m-n-1}) / a(\beta_+ - \beta_-)$. Here $\theta(n) \equiv 0$ for $n \leq 0$ and $\theta(n) \equiv 1$ for $n > 0$, and β_{\pm} are the solutions of the quadratic equation $a\beta^2 - (s+a+b)\beta + b = 0$. We define β_+ to be the solution that approaches s/a for large $|s|$. It can then be shown [10] that $|\beta_+| \geq |\beta_-|$ for all s . It follows that, at all s for which $|\beta_+| > 1$ (which is the case for all sufficiently large $|s|$), $\tilde{p}(1, x, s)$ must be given by $\tilde{p}(1, x, s) = (\beta_+)^{-n-1} / a$ in order for $\tilde{p}(m, x, s)$ to be bounded for large m . Hence $\tilde{G}(n, x, s)$, the Laplace transform of $G(n, x, t)$, is given by $\tilde{G}(n, x, s) = (\beta_+)^{-n-1}$. From Eq. (1), we then have the expression

$$P(\tau) = \sum_{n=0}^{\infty} \int_0^1 dx \int_{\gamma-i\infty}^{\gamma+i\infty} \frac{ds}{2\pi i} e^{s\tau} (1-b/a)(b/a)^n (\beta_+)^{-n-1} \quad (6)$$

for the desired quantity $P(\tau)$. Here γ is an arbitrary positive real number that is chosen large enough so that $|\beta_+| > 1$ for all s on the integration path. Performing the sum over n yields

$$P(\tau) = \int_0^1 dx \int_{\gamma-i\infty}^{\gamma+i\infty} \frac{ds}{2\pi i} e^{s\tau} (1-b/a)(\beta_+ - b/a)^{-1}. \quad (7)$$

Aside from the factor $e^{s\tau}$, the only s dependence of the integrand comes from β_+ . The only singularities of the integrand occur where β_+ itself has singularities, i.e., along a branch cut on the real axis between $s_- \equiv -(\sqrt{a} + \sqrt{b})^2$ and $s_+ \equiv -(\sqrt{a} - \sqrt{b})^2$. It is easy to show that this branch cut lies entirely to the left of the integration path in Eq. (7). [The factor $(\beta_+ - b/a)$ vanishes only if $\mu = \lambda$ and $x = 0$, and then only at $s = s_+ = 0$, which is the branch point at the right end of the branch cut.] The integral is performed by closing the contour at large $|s|$ in the left half plane. We now deform this contour into one that runs just above and below the branch cut and encloses it counterclockwise in standard fashion [11], yielding the following expression for $P(\tau)$:

$$P(\tau) = \int_0^1 dx \int_{s_-}^{s_+} \frac{ds}{2\pi i} e^{s\tau} (1-b/a) \frac{[(s-s_-)(s_+-s)]^{1/2}}{(-sb)}. \quad (8)$$

It remains only to extract the leading behavior of this double integral for asymptotically large τ . Note that, as τ gets large, the integral is dominated increasingly by values of s near the upper limit s_+ of the integration range. Writing $s = s_+ - u$, we have

$$P(\tau) = \int_0^1 dx \int_0^{s_+-s_-} du \frac{e^{(s_+-u)\tau} (a-b)[u(s_+-u-s_-)]^{1/2}}{2\pi b(u-s_+)}. \quad (9)$$

We now distinguish two cases.

(a) $\mu = \lambda$. Here $s_+ = -(\sqrt{a} - \sqrt{b})^2 = -\lambda x^2 / 4 + O(x^3)$, whereupon the largest values of the exponentials and hence the leading asymptotic behavior of the double integral come from the region where both x and u are near 0. Using $a-b = \lambda x$, we expand the integrand to leading order in x and u , obtaining, up to a multiplicative constant, $e^{-(\lambda x^2/4+u)\tau} x u^{1/2} / (Au + Bx^2)$, for positive constants A and B . Sending the upper limits on both the u and x integrals to infinity, we perform the rescalings $x = x' / \tau^{1/2}$ and $u = u' / \tau$ to yield $P(\tau) \sim \tau^{-3/2}$, the prefactor being a convergent double integral over x' and u' .

(b) $\mu > \lambda$. Here $a-b = \mu - \lambda + \lambda x$, and so approaches a nonzero constant as $x \rightarrow 0$, as does s_+ , which behaves like $s_+ \sim -[(\sqrt{\mu} - \sqrt{\lambda})^2 + vx + O(x^2)]$, where $v \equiv \sqrt{\lambda}(\sqrt{\mu} - \sqrt{\lambda})$. Thus the exponential factor in the integrand behaves as $e^{-(\sqrt{\mu} - \sqrt{\lambda})^2 \tau} e^{-[vx + u + O(x^2)]\tau}$, from which it follows that, again, the asymptotic behavior comes from x and u both near 0. All the

nonexponential factors of the integrand approach constants at $x=u=0$, except for the factor $u^{1/2}$. Sending the upper limits on the integrals to infinity and performing the rescalings $x=x'/\tau$ and $u=u'/\tau$, we find, for asymptotically large τ , $P(\tau) \sim \tau^{-5/2} e^{-\tau/\tau_0}$ [12], where $\tau_0 \equiv (\sqrt{\mu} - \sqrt{\lambda})^{-2}$.

It is instructive to compare this result with that for the continuous-task approximation of Ref. [6] [see Model A, case (2), $\lambda < \mu$]. There we obtained $P(\tau) \sim \tau^{-5/2} \exp(-\tau/\tau_0)$ where $1/\tau_0 = (\mu - \lambda)^2 / [4\mu(1 - \lambda)]$, μ and λ were probabilities per discrete time step, and τ_0 was the exponential time constant expressed as a number of time steps. If successive time steps are spaced Δt apart, the task execution and arrival rates are $\mu' = \mu/\Delta t$ and $\lambda' = \lambda/\Delta t$, and the exponential time constant is $\tau'_0 = \tau_0 \Delta t$. For the continuous-time limit, we take $\Delta t \rightarrow 0$, keeping μ' and λ' constant, and obtain $1/\tau'_0 = \mu'(1 - \eta)^2/4$ where $\eta \equiv \lambda'/\mu'$. By way of comparison, the present paper yields (affixing primes to keep the notation consistent) $1/\tau'_0 = (\sqrt{\mu'} - \sqrt{\lambda'})^2 = \mu'(1 - \sqrt{\eta})^2$. The critical point is at $\eta = 1$, where this case (b) changes over to case (a). In the limit $\eta \rightarrow 1$ from below, we have $\mu'(1 - \sqrt{\eta})^2 \rightarrow \mu'[(1 - \eta)^2/4 + O((1 - \eta)^3)]$, which agrees near the critical point, as it should, with the result from Ref. [6].

Finally, we consider the case $\mu < \lambda$, wherein tasks arrive more frequently than they are executed on average, produc-

ing a queue that grows in time t as $(\lambda - \mu)t$. As we pointed out in Ref. [6], however, tasks with priorities $x > x^* \equiv (\lambda - \mu)/\lambda$ are executed more frequently than they arrive on average, since their rate of arrival, $\lambda(1 - x)$, is less than μ . As in [6], therefore, the analysis of these tasks can be shown to be identical to that of case (a) above. In particular, the change of variable from x to w defined by $x = x^* + (1 - x^*)w$ maps the range $x^* < x < 1$ onto the range $0 < w < 1$, and transforms the quantities $a \equiv \mu$ and $b(x) \equiv \lambda(1 - x)$ in Eqs. (2) and (4) into $\tilde{a} = \tilde{\mu}$ and $\tilde{b}(w) = \lambda(1 - x^*)(1 - w) = \tilde{\lambda}(1 - w)$, respectively, where $\tilde{\mu} = \tilde{\lambda} = \mu$. Hence the problem of the tasks with $x > x^*$ maps precisely onto the original problem with $\mu = \lambda$, i.e., case (a) above, for which $P(\tau) \sim \tau^{-3/2}$.

Since, moreover, the total number of tasks with $x < x^*$ in the queue grows without bound as time progresses, the probability of executing such a task approaches 0 in the long-time limit. Asymptotically, therefore, all such tasks remain in the queue forever.

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- $< |\beta_-|$ if and only if $|\phi_+ + \phi_- - 2\phi| > \pi$, which is impossible. Also, $|\beta_+| = |\beta_-|$ only for s on the branch cut of β_+ and β_- , i.e., on the interval $[s_-, s_+]$ of the real s axis, where $s_{\pm} \equiv -(\sqrt{a} \mp \sqrt{b})^2$. Note also that, since $\beta_+ \beta_- = b/a \leq 1$, $|\beta_-| \leq 1$ for all s .
 [11] See, e.g., G. F. Carrier, M. Krook, and C. E. Pearson, *Functions of a Complex Variable: Theory and Technique* (McGraw-Hill, New York, 1966).
 [12] This result differs from the $P(\tau) \approx \tau^{-3/2} \exp(-\tau/\tau_0)$ form obtained numerically in Refs. [1–3]. Just as in Ref. [6], however, when λ is just slightly less than μ , making τ_0 large, there is an intermediate τ regime, $1 \ll \mu\tau \ll \mu\tau_0$, where the x^2 term of s_+ dominates the term of order x for almost the entire range of x . Thus the scaling of the x integral in Eq. (9) proceeds as for $\mu = \lambda$ [case (a)], yielding the result $P(\tau) \approx \tau^{-3/2} \exp(-\tau/\tau_0)$, valid in the intermediate τ regime, but not for asymptotically large τ .